

Week 8 Propensity Scores

Stat 209

Let $z=1,0$ T/C \underline{x} vector of covariates

propensity score $e(\underline{x}) = \Pr(z=1|\underline{x})$

scalar $\hat{e}(\underline{x})$

cond'l prob unit w/ vector \underline{x} observed cov. assigned to T ($z=1$)

Thm Balancing score $b(\underline{x})$ s.t. conditional distrib of \underline{x} given $b(\underline{x})$ same of treated and control units
 $\underline{x} \perp\!\!\!\perp z | b(\underline{x})$. Coarsest (low dimen) balancing score is propensity score. $\Pr(\underline{x}, z | e) = \Pr(\underline{x} | e) \Pr(z | e)$

Thm (result) Approx 90% reduction in bias for subclassifying at quintiles of population propensity score. $B_T = E(f(\underline{x}) | z=1) - E(f(\underline{x}) | z=0)$, B_S after stratification
 percent reduction in bias $100(1 - B_S/B_T) \approx 90\%$

- (i) The propensity score is a balancing score.
- (ii) Any score that is 'finer' than the propensity score is a balancing score; moreover, x is the finest balancing score and the propensity score is the coarsest.
- (iii) If treatment assignment is strongly ignorable given x , then it is strongly ignorable given any balancing score
- (iv) At any value of a balancing score, the difference between the treatment and control means is an unbiased estimate of the average treatment effect at that value of the balancing score if treatment assignment is strongly ignorable. Consequently, with strongly ignorable treatment assignment, pair matching on a balancing score, subclassification on a balancing score and covariance adjustment on a balancing score can all produce unbiased estimates of treatment effects.
- (v) Using sample estimates of balancing scores can produce sample balance on x .

Ros Rubin
 1983 Biometrika
 1984 JASA

Applications: Rubin Breast Cancer, Love (RR '84) CAD, Love Aspirin, Hansen SAT coaching, Substance Rosenbaum, Danish downers Abuse (UNC)

Robin AnnInt Medicine

Lalonde data

Lab 4 stratification

Table 3: Estimated 5-year Survival Rates for Node-Negative Patients in SEER from Tables 5 and 7 in U.S. GAO Report (1994).

AIM pub

Propensity Score

Subclass	Treatment	n	Estimate	n*	Estimate*
1	Breast Conservation	56	85.6%	54	88.8%
	Mastectomy	1,008	86.7%	966	90.5%
2	Breast Conservation	106	82.8%	102	86.0%
	Mastectomy	964	82.8%	917	87.7%
3	Breast Conservation	193	85.2%	184	89.4%
	Mastectomy	866	88.8%	841	91.4%
4	Breast Conservation	289	88.7%	279	92.0%
	Mastectomy	978	87.3%	742	91.5%
5	Breast Conservation	462	89.0%	453	90.7%
	Mastectomy	604	88.5%	589	90.7%

* omitting patients whose deaths were unrelated to cancer.

```
> table(propbin, treat)
      treat
propbin  0  1
(0,0.0401] 122  1
(0.0401,0.0872] 116  7
(0.0872,0.27] 101 21
(0.27,0.671] 53  71
(0.671,1] 37  85
> tapply(re78, list(propbin, treat), mean)
      0  1
(0,0.0401] 10467  0
(0.0401,0.0872] 5797 7919
(0.0872,0.27] 6043 9211
(0.27,0.671] 4977 5819
(0.671,1] 4666 6030
```

counts

means re78

The challenge of matching on the propensity to be coached

Histogram of propensity scores

Overlap??

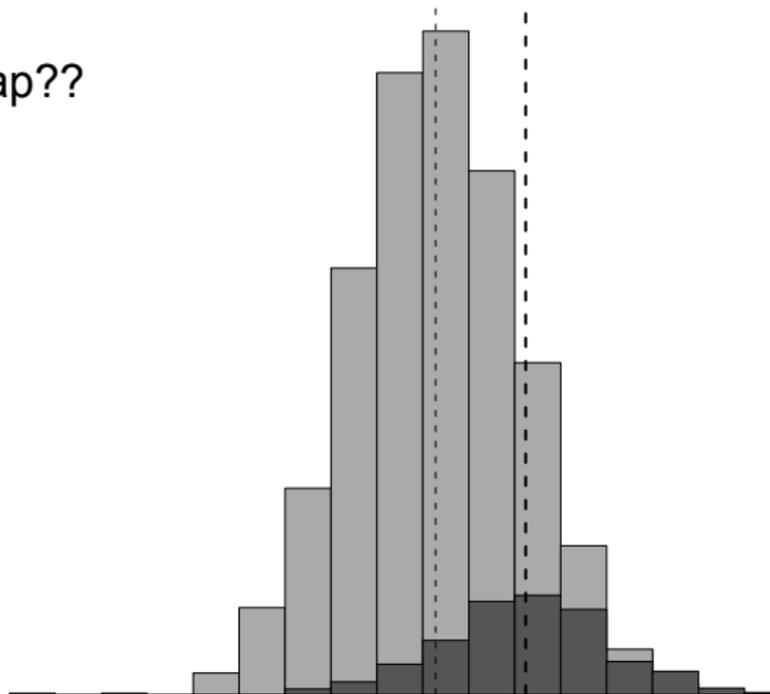


Table 1. Selected Pretreatment Variables

Variable	Range of values	Standardized bias	Percentage of sample
Math section of PSAT	20–43	–.1	18
	45–51	.1	17
	52–57	–.1	16
	58–80	.1	15
	Not taken	.1	34
Mean SAT at respondent's first-choice college	787–987	–.3	16
	988–1,060	–.2	16
	1,061–1,123	.1	16
	1,124–1,336	.3	16
Father's education	No response	.0	36
	High school	–.4	40
	A.A. or B.A.	–.1	26
	Graduate	.4	25
Average math grade	No response	.2	9
	"Excellent"	.1	35
	"Good"–"fail"	–.1	59
Foreign language years taken	No response	.1	6
	0–2	–.3	64
	3–4	.3	27
	No response	.1	9

well as scores on previous SAT–I or PSAT tests and their answers to the Student Descriptive Questionnaire (SDQ), which all SAT–I registrants are asked to complete. By their responses to questions about extracurricular SAT preparation, respondents split into a treated and a control group, and the data describe the results of a classical quasiexperiment (Campbell and Stanley 1966).

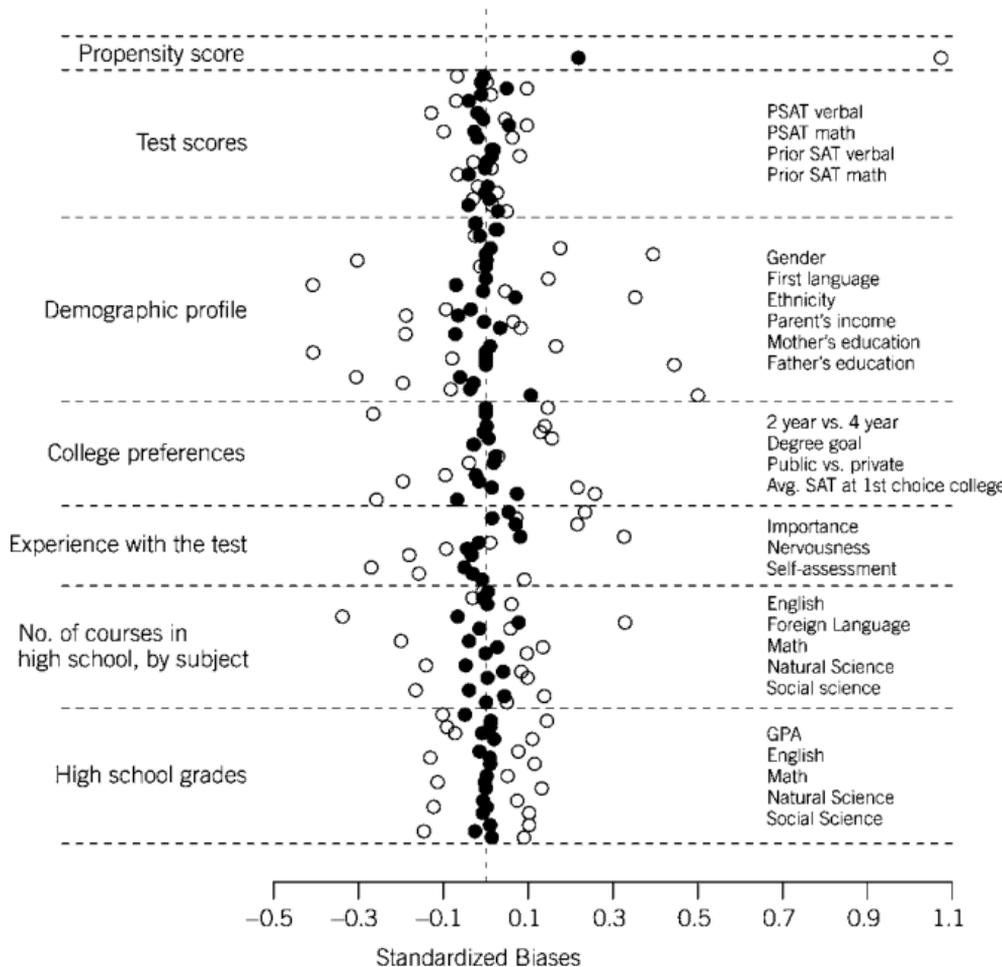
Nineteen in twenty of the survey respondents actually took the spring 1996 or fall 1995 exam for which they had registered. The analysis given below restricts itself to these 3,994 students, using the corresponding SAT scores as outcome measures. Thus the record gives coaching status and SAT outcomes for all students in the sample to be analyzed; among the additional measures, each available for some fraction of the students, are pretest scores, racial and socioeconomic indicators, various data about their academic preparation, and responses to a survey item that, by eliciting students' first choices in colleges, recovered an unusually discriminating measure of students' educational aspirations. In all, there are 27 pretreatment variables.

The coached and uncoached groups differ appreciably in these recorded measures—as do high and low scorers on the SAT. Table 1 offers some illustration of this, giving overall incidences of various covariate attributes and comparing their relative incidences in the coached and uncoached groups.

(The statistic here used to effect these comparisons is the *standardized bias*, given for a variable v by $(\bar{v}_t - \bar{v}_c)/s_p$, where \bar{v}_t and \bar{v}_c are the average values of v in the treatment and control groups, respectively, and s_p^2 is the pooled within-group variance in v .) Yet the table shows only five covariates; the analysis must address biases on all 27 of them.

1.2 Missing and Misleading Data in Regression and in Subclassification

In regression-based adjustment, the simplest way to handle missing data on a covariate is to reject cases without complete information. In adjustment based on matching or stratification,



Multivariate Matching with the Propensity Score

- Match subjects so that they balance on multiple covariates using one scalar score.
- Goal: **Emulate a RCT in matching, then use standard analyses to compare matched sets.**
- Design: Treated subjects matched to people who didn't receive treatment but who had similar propensity to receive treatment (match the treated to untreated "clones").

Aspirin Use and Mortality

- **6174 consecutive** adults at CCF undergoing stress echocardiography for evaluation of known or suspected coronary disease.
- **2310 (37%)** were taking aspirin (treatment).
- Main Outcome: all-cause mortality
- Median follow-up: 3.1 years
- Univariate Analysis: 4.5% of aspirin patients died, and 4.5% of non-aspirin patients died...
- **Unadjusted Hazard Ratio: 1.08 (0.85, 1.39)**

Gum et al. (2001) <http://www.ncbi.nlm.nih.gov/pubmed/11559263>

JAMA. 2001 Sep 12;286(10):1187-94. <http://jama.jamanetwork.com/article.aspx?articleid=194177>
Aspirin use and all-cause mortality among patients being evaluated for known or suspected coronary artery disease: A propensity analysis. Gum PA1, Thamarasan M, Watanabe J, Blackstone EH, Lauer MS.

Propensity Score Model for Aspirin Use

- Logistic Regression predicting aspirin use
- **31 covariates included in the model:**
 - Demographics, Clinical history, Medication use
 - Cardiovascular assessment and Exercise capacity
- Estimated propensity scores for aspirin use range from .03 to .98
 - ROC Area shows good discrimination (C = .83)
- But does the propensity score model work?
- Are the covariates balanced?

Baseline Characteristics By Aspirin Use (in %) (before matching)

Variable	Aspirin (n = 2310)	No Aspirin (n = 3864)	P value
Men	77.0	56.1	< .001
Clinical history: diabetes	16.8	11.2	< .001
hypertension	53.0	40.6	< .001
prior coronary artery disease	69.7	20.1	< .001
congestive heart failure	5.5	4.6	.12
Medication use: Beta-blocker	35.1	14.2	< .001
ACE inhibitor	13.0	11.4	< .001

- Baseline characteristics appear very dissimilar: 25 of 31 covariates have $p < .001$, **28 of 31 have $p < .05$.**
- Aspirin user covariates indicate higher mortality risk.

Propensity Matcher Results

ID	Treated?	Propensity	Linear Propensity	Match?	Partner ID
1	1	0.2	-1.386	No	-999
2	1	0.3	-0.847	Yes	8
3	1	0.4	-0.405	Yes	10
4	1	0.6	0.405	No	-999
5	1	0.7	0.847	No	-999

logit

SE (Linear Propensity):	0.1829
x % Selected:	0.6
x % of SE:	0.1097

Baseline Characteristics By Aspirin Use [%] (after matching)

Variable	Aspirin (n = 1351)	No Aspirin (n = 1351)	P value
Men	70.4	72.1	.33
Clinical history: diabetes	15.0	15.3	.83
hypertension	50.3	51.7	.46
prior coronary artery disease	48.3	48.8	.79
congestive heart failure	5.8	6.6	.43
Medication use: Beta-blocker	26.1	26.5	.79
ACE inhibitor	15.5	15.8	.79

- Baseline characteristics similar in matched users and non-users.
- 30 of 31 covariates show NS difference between matched users and non-users. [Peak exercise capacity for men is p = .01]

Using Standardized Differences to Measure Covariate Balance

- Standardized Differences are appropriate summaries of Covariate Balance for both Continuous and Categorical Variables

$$d = \frac{100(\bar{x}_{Treatment} - \bar{x}_{Control})}{\sqrt{\frac{s_{Treatment}^2 + s_{Control}^2}{2}}} \text{ for continuous variables}$$

$$d = \frac{100(p_{Treatment} - p_{Control})}{\sqrt{\frac{p_T(1-p_T) + p_C(1-p_C)}{2}}} \text{ for binary variables}$$

|Standardized Differences| > 10% Indicate Serious Imbalance

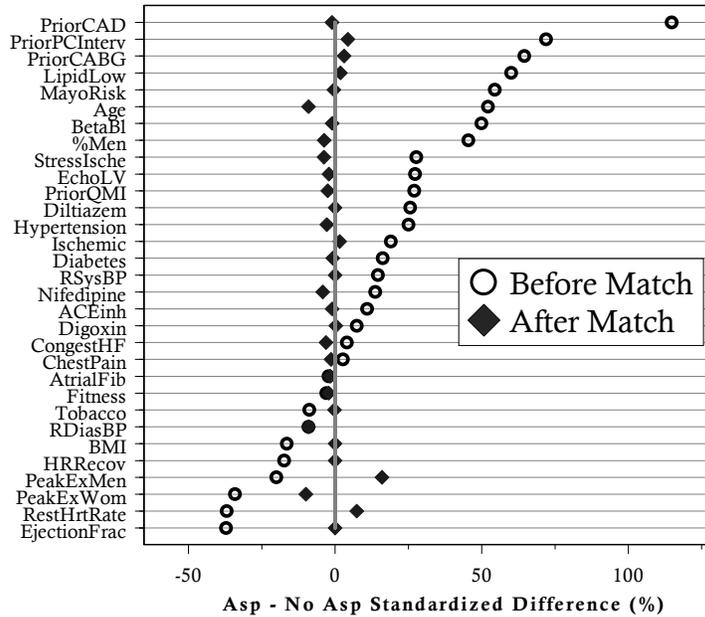
Before Match:

- 811/2310 (35.1%) Aspirin users used β -blockers
- 550/3864 (14.2%) non-Aspirin users used β -blockers
- Standardized Difference is 49.9%
- P value for difference is < .001

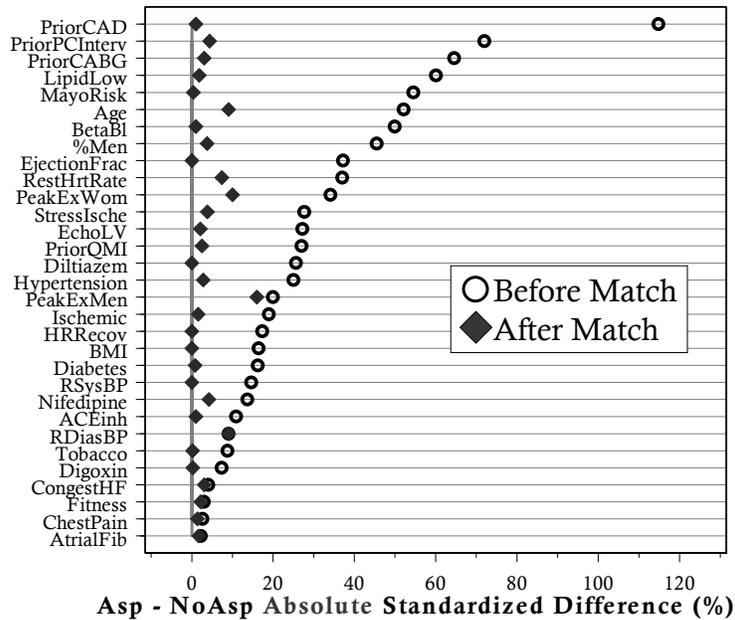
After Match:

- 352/1351 (26.1%) Aspirin users used β -blockers
- 358/1351 (26.5%) non-Aspirin users used β -blockers
- Standardized Difference is –1.0%
- P value for difference is .79

Covariate Balance for Aspirin Study



Absolute Standardized Differences



In other words, x and t are, necessarily, conditionally independent variables given the propensity score, $p = \Pr(t = 1 | x)$. This is really a very simple theorem in statistics / probability that requires only rather weak assumptions. In fact, the real “problem” in applications is simply that the functional form of the true PS is usually unknown and, thus, needs to be estimated from data!

When the conditional distributions of baseline patient characteristics and treatment choices “fail to factor” as dictated by the fundamental theorem of PS, [2], this is rightfully interpreted as clear evidence that one’s estimates of the unknown, true PSs are not even approximately correct.

4. Case Study Example: Effects of Abciximab use on both Survival and Cardiac Billing.

The (numerical and graphical) output illustrations provided here use the data from Kereiakes et al. (2000). The corresponding command file and data are distributed along with the UPS and SPS “R” functions in the files ABCIXini.R and ABCIX.CSV. In this prospective study, **outcomes variables (survival and cardiac related costs)** were collected via follow-up for at least 6 months on **996 PCI** patients treated at the **Lindner Center**, Christ Hospital, Cincinnati. **Rather than randomize patients to treatment, the Lindner interventionists practiced “evidence based medicine” in choosing between either augmenting or not augmenting their “usual care” for Percutaneous Coronary Intervention (PCI) with abciximab (Reopro®), a relatively expensive IIb/IIIa cascade inhibitor.** Ability-to-pay was not a factor in this treatment choice in the sense that Lindner interventionists had access to “research use” abciximab.

Our objective in this “R user’s manual” documentation for the UPS and SPS functions is not to fully discuss and illustrate all aspects of the abciximab case study. Rather, we simply wish to illustrate some example UPS and SPS function invocations as well as the tabular and/or graphical output that results. Readers interested in reading more about UPS and SPS analyses using the abciximab case study are referred to my “white paper,” Obenchain(2006a).

Variables in the Kereiakes et al (2000) **Abciximab / Lindner Data:**

Description	Name	Values
Life Years Preserved = 0 if died within 6 Months or 11.6 Years given Survival for at least 6 Months	lifepres	Either 0 or 11.6 Years
Total Cardiac Related Billing within 12 Months of Patient’s Initial PCI at Lindner Center	cardbill	\$2,216 to \$178,534 in 1997 US Dollars
Was “Usual PCI Care” augmented with Abciximab?	abcix	0 => No, 1 => Yes
Was a Stent (anti-collapse device) Deployed?	stent	0 => No, 1 => Yes
Patient Height in Centimeters	height	108 cm to 196 cm
Patient Gender	female	0 => No, 1 => Yes
Was the patient Diabetic?	diabetic	0 => No, 1 => Yes
Had patient suffered an Acute Myocardial Infarction within the Last Seven Days?	acutemi	0 => No, 1 => Yes
Left Ventricular Ejection Fraction	ejecfrac	0% to 90%
Number of Vessels involved in first PCI procedure	veslproc	0 to 5

without reference to strata.

2. Lindner dataset

We make use of the observational data frame `lindner`, first provided with the R package **USPS**, by [Obenchain \(2007\)](#). The `lindner` data contain data on 996 patients treated at the Lindner Center, Christ Hospital, Cincinnati in 1997. Patients received a Percutaneous Coronary Intervention (PCI). The data consists of 10 variables. Two are outcome: `lifepres` ranges over two values, 11.4 or 0 depending on whether patients survived to six months post treatment or did not survive to six months (respectively), where 11.4 is the mean expected preserved life years for those patients who survive to six months. Secondly, `cardbill` contains the costs in 1998 dollars for the first six months (or less if patient did not survive) after treatment with the drug abciximab. For simplicity we confine our analysis to `cardbill`, focusing on the logarithm of this variate. The treatment variable is `abcix`, where 0 indicates standard PCI treatment and 1 indicates standard PCI treatment and additional treatment in some form with abciximab. Covariates include `acutemi`, 1 indicating a recent acute myocardial infarction and 0 not; `ejecfrac` for the left ventricle ejection fraction, a percentage from 0 to 90; `ves1proc` giving the number of vessels (0 to 5) involved in the initial PCI; `stent` with 1 indicating coronary stent inserted, 0 not; `diabetic` where 1 indicates that the patient has been diagnosed with diabetes, 0 not; `height` in centimeters and `female` coding the sex of the patient, 1 for female, 0 male.

3. Estimation of propensity scores, production of strata

We estimate propensity scores in two ways: via logistic regression and a recursively partitioned tree.

```
R> data("lindner")
R> attach(lindner)
R> lindner.log <- glm(abcix ~ stent + height + female + diabetic +
+   acutemi + ejecfrac + ves1proc, data = lindner, family = binomial)
R> ps <- lindner.log$fitted
R> lindner.s5 <- cut(ps, quantile(ps, seq(0, 1, 1/5)),
+   include.lowest = TRUE, labels = FALSE)
R> lindner.s10 <- cut(ps, quantile(ps, seq(0, 1, 1/10)),
+   include.lowest = TRUE, labels = FALSE)
```

In `lindner.s5` we produced five strata of roughly equal size, ten strata in `lindner.s10`. Alternatively, propensity scores can be estimated and strata defined using a recursively partitioned tree. This generally provides strata that generally differ in their sizes. We use the R function `rpart` from the package of the same name by [Therneau and Atkinson \(2002\)](#) to generate a tree with six bins or strata; the tree is not reproduced here.

```
R> library("rpart")
R> lindner.rpart <- rpart(abcix ~ stent + height + female + diabetic +
```

Examples

```
#Note reordering of columns, binary factor and numeric column are unchanged.
f2 <- factor(sample(c(0, 1), 20, replace = TRUE))
f4 <- factor(sample(c("a", "b", "c", "d"), 20, replace = TRUE))
cv <- rnorm(20)
X <- data.frame(f2, f4, cv)
cv.trans.psa(X)
#
f2 <- factor(sample(c('c', 'C'), 20, replace = TRUE))
f4 <- factor(sample(c("b", "A", "d", "CC"), 20, replace = TRUE))
cv <- rnorm(20)
X <- data.frame(f2, f4, cv)
cv.trans.psa(X)
```

lindner

Data on 996 initial Percutaneous Coronary Interventions (PCIs) performed in 1997 at the Lindner Center, Christ Hospital, Cincinnati.

Description

Data from an observational study of 996 patients receiving a PCI at Ohio Heart Health in 1997 and followed for at least 6 months by the staff of the Lindner Center. This is a landmark dataset in the literature on propensity score adjustment for treatment selection bias due to practice of evidence based medicine; patients receiving abciximab tended to be more severely diseased than those who did not receive a IIb/IIIa cascade blocker.

Usage

```
data(lindner)
```

Format

A data frame with 996 observations on the following 10 variables, no NAs.

lifepres Mean life years preserved due to survival for at least 6 months following PCI; numeric value of either 11.4 or 0.

cardbill Cardiac related costs incurred within 6 months of patient's initial PCI; numeric value in 1998 dollars; costs were truncated by death for the 26 patients with lifepres == 0.

abcix Numeric treatment selection indicator; 0 implies usual PCI care alone; 1 implies usual PCI care deliberately augmented by either planned or rescue treatment with abciximab.

stent Coronary stent deployment; numeric, with 1 meaning YES and 0 meaning NO.

height Height in centimeters; numeric integer from 108 to 196.

female Female gender; numeric, with 1 meaning YES and 0 meaning NO.

```
##### Week 2 Computing Corner, Rogosa R-session
```

```
> library(PSAgraphics)
```

```
> data(lindner)
```

```
> attach(lindner)
```

```
> dim(lindner)
```

```
[1] 996 10
```

```
> head(lindner)
```

```
  lifepres  cardbill  abcix  stent  height  female  diabetic  acutemi  ejecfrac  veslproc
1      0.0    14301      1      0    163      1      1      0      56      1
2     11.6    3563      1      0    168      0      0      0      56      1
3     11.6    4694      1      0    188      0      0      0      50      1
4     11.6    7366      1      0    175      0      1      0      50      1
5     11.6    8247      1      0    168      1      0      0      55      1
6     11.6    8319      1      0    178      0      0      0      50      1
```

```
> str(lindner)
```

```
'data.frame': 996 obs. of 10 variables:
```

```
$ lifepres: num 0 11.6 11.6 11.6 11.6 11.6 11.6 11.6 11.6 11.6 ...
```

```
$ cardbill: int 14301 3563 4694 7366 8247 8319 8410 8517 8763 8823 ...
```

```
$ abcix : int 1 1 1 1 1 1 1 1 1 1 ...
```

```
$ stent : int 0 0 0 0 0 0 0 0 0 0 ...
```

```
$ height : int 163 168 188 175 168 178 185 173 152 180 ...
```

```
$ female : int 1 0 0 0 1 0 0 1 1 0 ...
```

```
$ diabetic: int 1 0 0 1 0 0 0 0 0 0 ...
```

```
$ acutemi : int 0 0 0 0 0 0 0 0 0 0 ...
```

```
$ ejecfrac: int 56 56 50 50 55 50 58 30 60 60 ...
```

```
$ veslproc: int 1 1 1 1 1 1 1 1 1 1 ...
```

```
> table(abcix)
```

```
abcix
 0  1
298 698
```

```
##### look at outcomes
```

```
> tapply(cardbill, abcix, mean)
```

```
 0      1
14614.22 16126.68
```

```
> tapply(log(cardbill), abcix, mean) # analyses done in log scale
```

```
 0      1
9.398158 9.581579
```

```
> t.test(log(cardbill) ~ abcix) # treatment leads to higher bills?
```

```
Welch Two Sample t-test
```

```
data: log(cardbill) by abcix
```

```
t = -5.2317, df = 461, p-value = 2.554e-07
```

```
alternative hypothesis: true difference in means is not equal to 0
```

```
95 percent confidence interval:
```

```
-0.2523168 -0.1145249
```

```
sample estimates:
```

```
mean in group 0 mean in group 1
 9.398158      9.581579
```

```
> tapply(lifepres, abcix, mean)
```

```
 0      1
11.01611 11.41719
```

```
> table(lifepres, abcix)
```

```
      abcix
lifepres  0  1
 0      15  11
11.6    283 687
```

```
> chisq.test(lifepres, abcix)
```

```
Pearson's Chi-squared test with Yates' continuity correction
```

```
data: lifepres and abcix
```

X-squared = 8.5077, df = 1, p-value = 0.003536

```
> prop.table(table(lifepres, abcix),2) # look at relative risk
      abcix
lifepres  0      1
0      0.05033557 0.01575931
11.6  0.94966443 0.98424069
```

```
> library(MatchIt) ## try full matching
```

```
> m2full = matchit(abcix ~ stent + height + female + diabetic + acutemi + ejecfrac + veslproc,
                  data = lindner, method = "full")
```

Loading required package: optmatch

Loading required package: survival

```
> detach(lindner)
```

```
> m2full.dat = match.data(m2full)
```

```
> attach(m2full.dat)
```

```
> boxplot(distance ~ abcix) # propensity score overlap
> # can also do overlapping histogram as in week 1
```

```
> summary(m2full) # check balance improvement
```

```
Call: matchit(formula = abcix ~ stent + height + female + diabetic +
              acutemi + ejecfrac + veslproc, data = lindner, method = "full")
```

Summary of balance for all data:

	Means Treated	Means Control	Mean Diff	eQQ Med	eQQ Mean	eQQ Max
distance	0.7265	0.6406	0.0859	0.0814	0.0852	0.1209
stent	0.7049	0.5839	0.1210	0.0000	0.1208	1.0000
height	171.4427	171.4463	-0.0036	0.0000	0.5638	20.0000
female	0.3309	0.3859	-0.0550	0.0000	0.0537	1.0000
diabetic	0.2049	0.2685	-0.0636	0.0000	0.0638	1.0000
acutemi	0.1791	0.0604	0.1187	0.0000	0.1174	1.0000
ejecfrac	50.4026	52.2886	-1.8860	1.0000	2.0503	20.0000
veslproc	1.4628	1.2047	0.2581	0.0000	0.2651	1.0000

Summary of balance for matched data:

	Means Treated	Means Control	Mean Diff	eQQ Med	eQQ Mean	eQQ Max
distance	0.7265	0.7262	0.0003	0.0068	0.0077	0.0798
stent	0.7049	0.7465	-0.0416	0.0000	0.0248	1.0000
height	171.4427	171.6093	-0.1666	0.0000	0.9548	15.0000
female	0.3309	0.3016	0.0293	0.0000	0.0244	1.0000
diabetic	0.2049	0.2210	-0.0162	0.0000	0.0068	1.0000
acutemi	0.1791	0.1605	0.0186	0.0000	0.0300	1.0000
ejecfrac	50.4026	50.9846	-0.5821	0.0000	0.9524	20.0000
veslproc	1.4628	1.4616	0.0012	0.0000	0.0324	1.0000

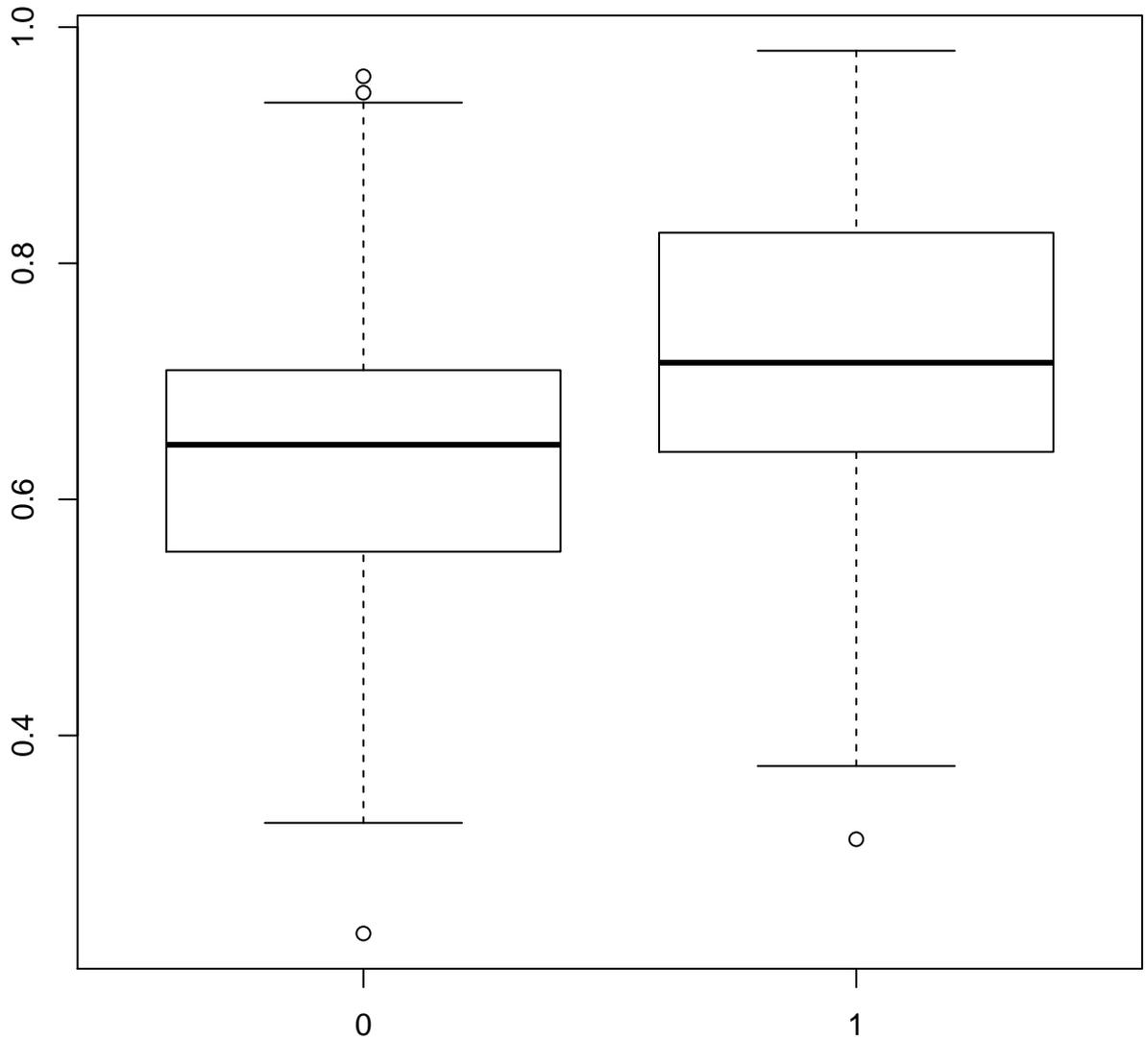
Percent Balance Improvement:

	Mean Diff	eQQ Med	eQQ Mean	eQQ Max
distance	99.5977	91.6649	91.0174	33.9503
stent	65.6178	0.0000	79.4711	0.0000
height	-4508.6968	0.0000	-69.3633	25.0000
female	46.6926	0.0000	54.5550	0.0000
diabetic	74.5771	0.0000	89.3347	0.0000
acutemi	84.3069	0.0000	74.4571	0.0000
ejecfrac	69.1376	100.0000	53.5491	0.0000
veslproc	99.5373	0.0000	87.7782	0.0000

Sample sizes:

	Control	Treated
All	298	698
Matched	298	698
Unmatched	0	0
Discarded	0	0

Propensity score boxplot



```
> summary(m2full, standardize = T) # look at standardized mean diffs < .1
Call: matchit(formula = abcix ~ stent + height + female + diabetic +
  acutemi + ejecfrac + veslproc, data = lindner, method = "full")
```

Summary of balance for all data:

	Means Treated	Means Control	Std. Mean Diff.	eCDF Med	eCDF Mean	eCDF Max
distance	0.7265	0.6406	0.6609	0.1777	0.1714	0.2760
stent	0.7049	0.5839	0.2651	0.0605	0.0605	0.1210
height	171.4427	171.4463	-0.0003	0.0060	0.0079	0.0250
female	0.3309	0.3859	-0.1167	0.0275	0.0275	0.0550
diabetic	0.2049	0.2685	-0.1574	0.0318	0.0318	0.0636
acutemi	0.1791	0.0604	0.3093	0.0593	0.0593	0.1187
ejecfrac	50.4026	52.2886	-0.1810	0.0114	0.0356	0.1138
veslproc	1.4628	1.2047	0.3654	0.0091	0.0433	0.1884

Summary of balance for matched data:

	Means Treated	Means Control	Std. Mean Diff.	eCDF Med	eCDF Mean	eCDF Max
distance	0.7265	0.7262	0.0027	0.0128	0.0123	0.0300
stent	0.7049	0.7465	-0.0911	0.0240	0.0240	0.0480
height	171.4427	171.6093	-0.0156	0.0128	0.0144	0.0444
female	0.3309	0.3016	0.0622	0.0094	0.0094	0.0188
diabetic	0.2049	0.2210	-0.0400	0.0044	0.0044	0.0088
acutemi	0.1791	0.1605	0.0485	0.0080	0.0080	0.0160
ejecfrac	50.4026	50.9846	-0.0559	0.0080	0.0109	0.0604
veslproc	1.4628	1.4616	0.0017	0.0044	0.0071	0.0196

Percent Balance Improvement:

	Std. Mean Diff.	eCDF Med	eCDF Mean	eCDF Max
distance	99.5977	92.7960	92.8062	89.1302
stent	65.6178	60.3235	60.3235	60.3235
height	-4508.6968	-112.9961	-81.7392	-77.8086
female	46.6926	65.7936	65.7936	65.7936
diabetic	74.5771	86.1603	86.1603	86.1603
acutemi	84.3069	86.5184	86.5184	86.5184
ejecfrac	69.1376	29.9650	69.4106	46.9404
veslproc	99.5373	51.4731	83.5345	89.5977

Sample sizes:

	Control	Treated
All	298	698
Matched	298	698
Unmatched	0	0
Discarded	0	0

```
> plot(summary(m2full, standardize = T)) # see plot, balance improvement
```

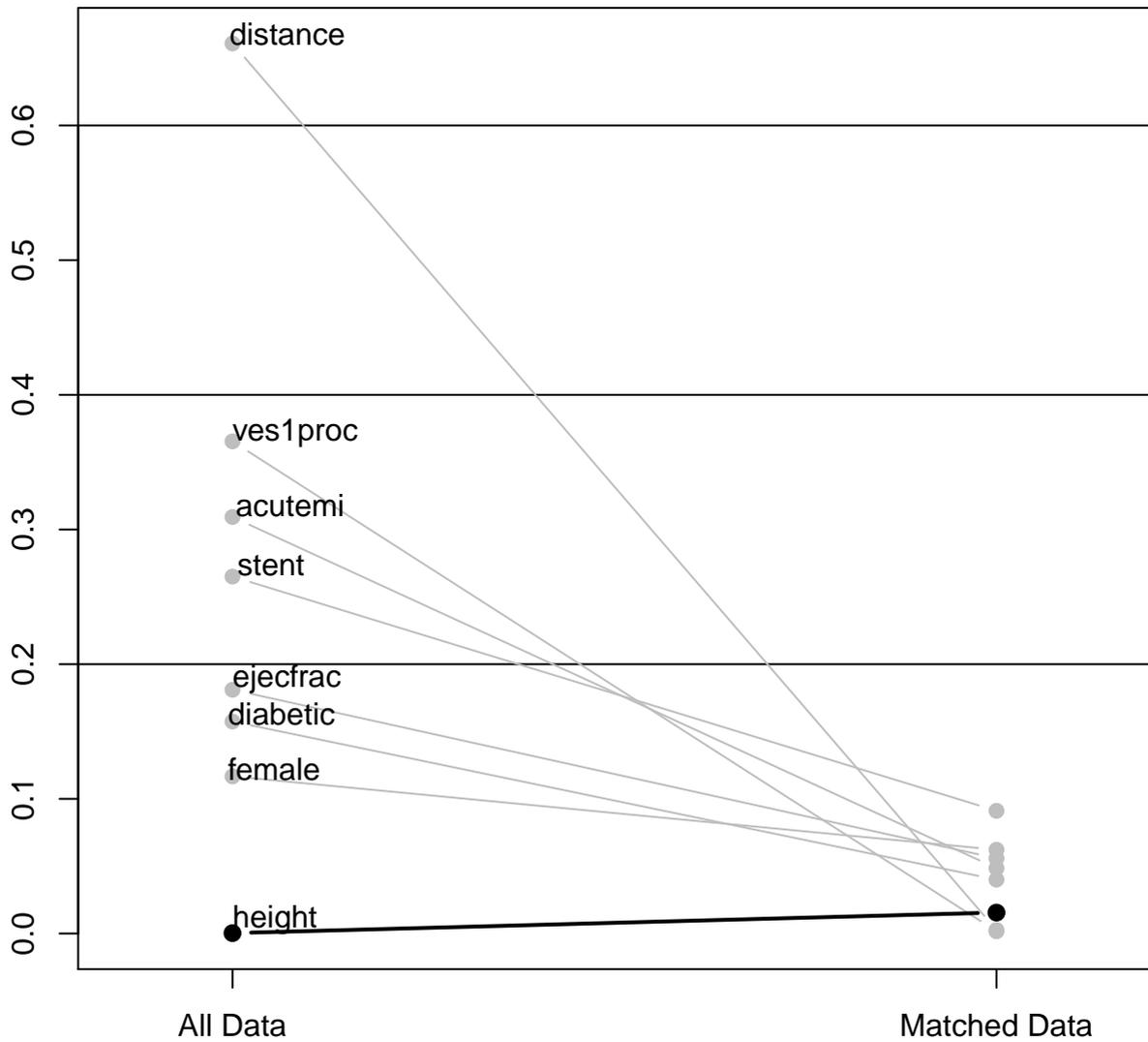
```
> head(m2full.dat)
```

	lifepres	cardbill	abcix	stent	height	female	diabetic	acutemi	ejecfrac	veslproc	distance	weights	subc
1	0.0	14301	1	0	163	1	1	0	56	1	0.4079170	1	
2	11.6	3563	1	0	168	0	0	0	56	1	0.5784602	1	
3	11.6	4694	1	0	188	0	0	0	50	1	0.5244469	1	
4	11.6	7366	1	0	175	0	1	0	50	1	0.4727311	1	
5	11.6	8247	1	0	168	1	0	0	55	1	0.4930466	1	
6	11.6	8319	1	0	178	0	0	0	50	1	0.5625524	1	

```
> table(subclass) #267 subclasses (698 treated) mostly small #> table(subclass, abcix) breakdown
```

subclass	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26
3	4	2	2	5	8	3	2	4	11	5	2	2	2	3	3	5	2	3	4	6	3	3	7	2	9	
33	34	35	36	37	38	39	40	41	42	43	44	45	46	47	48	49	50	51	52	53	54	55	56	57	58	
8	2	3	6	2	3	5	2	2	3	2	4	2	4	2	6	2	5	4	5	4	2	3	2	2	9	
65	66	67	68	69	70	71	72	73	74	75	76	77	78	79	80	81	82	83	84	85	86	87	88	89	90	
2	3	4	2	2	2	3	2	4	2	2	6	2	2	2	2	2	5	6	2	5	2	3	2	3		
97	98	99	100	101	102	103	104	105	106	107	108	109	110	111	112	113	114	115	116	117	118	119	120	121	122	
2	2	2	4	2	2	2	4	3	2	4	2	6	4	3	17	2	3	5	2	3	5	2	5	7	10	

Absolute Standardized Diff in Means



```

129 130 131 132 133 134 135 136 137 138 139 140 141 142 143 144 145 146 147 148 149 150 151 152 153 154
  17  4  8  8  4  4  6  2  2  3  3  3  4  5  2  2  6  6  4  7  10  3  2  2  8  2
161 162 163 164 165 166 167 168 169 170 171 172 173 174 175 176 177 178 179 180 181 182 183 184 185 186
  3  2  2  2  4  3  3  2  2  3  3  2  2  2  2  6  2  3  2  3  18  8  17  13  10  5
193 194 195 196 197 198 199 200 201 202 203 204 205 206 207 208 209 210 211 212 213 214 215 216 217 218
  2  2  10  2  15  2  4  9  3  2  7  3  4  4  3  2  5  2  2  4  7  2  2  2  4  2
225 226 227 228 229 230 231 232 233 234 235 236 237 238 239 240 241 242 243 244 245 246 247 248 249 250
  2  2  2  2  2  5  2  4  2  2  2  3  7  2  2  2  2  3  2  2  2  2  2  3  2  2
257 258 259 260 261 262 263 264 265 266 267
  2  2  2  2  2  2  2  2  5  2  2

```

```
> fivenum(table(subclass))
```

```
subclass
 3 154 266 147 190
 2  2  2  4  18
```

```
> library(lme4) compare_outcomes_over_subclasses, log(cardbill) outcome
```

```
> mfullL.lmer = lmer(log(cardbill) ~ abcix + (1 + abcix|subclass), data = m2full.dat)
```

```
> summary(mfullL.lmer)
```

```
Linear mixed model fit by REML ['lmerMod']
```

```
Formula: log(cardbill) ~ abcix + (1 + abcix | subclass) Data: m2full.dat
```

```
REML criterion at convergence: 1268.2
```

```
Scaled residuals:
```

```
      Min       1Q   Median       3Q      Max
-3.2325 -0.6183 -0.2411  0.4093  4.2623
```

```
Random effects:
```

```
Groups   Name             Variance Std.Dev. Corr
subclass (Intercept)  0.09744  0.3122
          abcix         0.06809  0.2609  -0.98
Residual                0.18034  0.4247
```

```
Number of obs: 996, groups: subclass, 267
```

```
Fixed effects:
```

```
              Estimate Std. Error t value
(Intercept)  9.40576    0.03159  297.71
abcix         0.16440    0.03417   4.81
```

```
Correlation of Fixed Effects:
```

```
(Intr)
abcix -0.866
```

```
> confint(mfullL.lmer)
```

```
Computing profile confidence intervals ...
```

```
          2.5 %      97.5 %
.sig01    0.22608548  0.3882918
.sig02   -1.00000000 -0.8884949
.sig03    0.15108346  0.3579166
.sigma    0.40090446  0.4492170
(Intercept) 9.34378346  9.4678932
abcix      0.09694981  0.2327080
```

```
> exp(confint(mfullL.lmer))
```

```
Computing profile confidence intervals ...
```

```
          2.5 %      97.5 %
.sig01  1.253683e+00  1.474460e+00
.sig02  3.678794e-01  4.112743e-01
.sig03  1.163094e+00  1.430346e+00
.sigma  1.493175e+00  1.567085e+00
(Intercept) 1.142756e+04  1.293760e+04
abcix      1.101805e+00  1.262013e+00
```

```
> # note mfull.lmer = lmer(cardbill ~ abcix + (1 + abcix|subclass), data = m2full.dat)
```

```
# gives non-sig result CI: abcix -497.3218 3009.1425564
```

```
#### treat lifepres as a 0,1 outcome
```

```

> mfull.glmer = glmer(as.factor(lifepres) ~ abcix + (1 + abcix|subclass),
                      family = binomial, data = m2full.dat)
> summary(mfull.glmer) # in log-odds metric
Generalized linear mixed model fit by maximum likelihood (Laplace Approximation) ['glmerMod']
Family: binomial ( logit )
Formula: as.factor(lifepres) ~ abcix + (1 + abcix | subclass)    Data: m2full.dat

```

AIC	BIC	logLik	deviance	df.resid
233.8	258.3	-111.9	223.8	991

```

Scaled residuals:
  Min      1Q  Median      3Q      Max
-4.3416  0.0110  0.0112  0.2300  0.8194

```

```

Random effects:
 Groups   Name      Variance Std.Dev. Corr
subclass (Intercept)  0.003532  0.05943
          abcix      42.902893  6.55003  1.00
Number of obs: 996, groups: subclass, 267

```

```

Fixed effects:
              Estimate Std. Error z value Pr(>|z|)
(Intercept)   2.9389     0.2659  11.053 < 2e-16 ***
abcix         6.0279     0.8381   7.192 6.39e-13 ***
---

```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Correlation of Fixed Effects:
  (Intr)
abcix -0.305
> exp(fixef(mfull.glmer)) #odds of survival
(Intercept)      abcix
 18.89487      414.82764

```

```

##### alternative, matching by regression interpolation, ancova vs blocking
> # propensity ancova, propensity score (distance) as covariate
> pancL = lm(log(cardbill) ~ abcix + distance, data =m2full.dat)
> summary(pancL)
Call: lm(formula = log(cardbill) ~ abcix + distance, data = m2full.dat)

```

```

Residuals:
  Min      1Q  Median      3Q      Max
-1.7137 -0.2979 -0.1257  0.2069  2.7026

```

```

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)   9.17752     0.07759  118.286 < 2e-16 ***
abcix         0.15384     0.03322   4.630 4.14e-06 ***
distance      0.34443     0.11378   3.027 0.00253 **
---

```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Residual standard error: 0.4589 on 993 degrees of freedom
Multiple R-squared:  0.04107, Adjusted R-squared:  0.03914
F-statistic: 21.26 on 2 and 993 DF, p-value: 9.075e-10

```

```

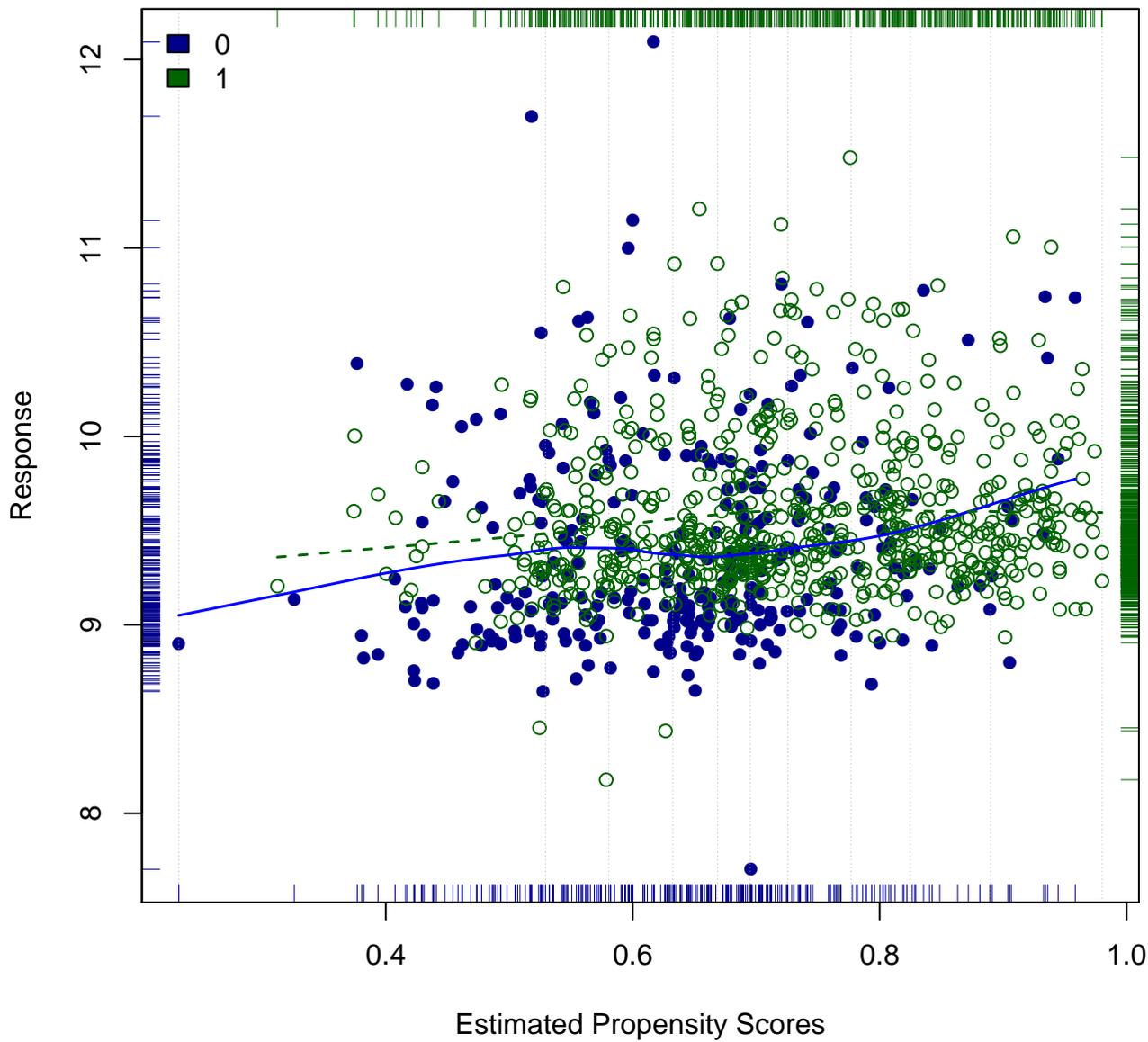
> exp(confint(pancL)) # almost same result as fullmatch lmer
              2.5 %      97.5 %
(Intercept) 8310.388799 11268.542739
abcix        1.092686   1.244877
distance     1.128783   1.764223

```

```

> pLancL = loess.psa(log(cardbill), abcix, distance) # use smoothers rather than straight-line
> pLancL # similar result; loess.psa from PSAGraphics also generates plot
$ATE
[1] 0.1279487

```



```
$se.wtd
[1] 0.04151743
```

```
$CI95
[1] 0.04491388 0.21098361
```

```
# compare a smoother ancova package fANCOVA
```

```
> install.packages("fANCOVA")
> library(fANCOVA)
> pLanc2L = loess.ancova(distance, log(cardbill), abcix)
> pLanc2L
$linear.fit
```

```
          [,1]
(Intercept) 9.3863321
group1      0.1522632
```

```
$smooth.fit
```

```
Call:
```

```
loess(formula = lm.res ~ x, span = span1, degree = degree, family = family)
```

```
Number of Observations: 996
```

```
Equivalent Number of Parameters: 5.28
```

```
Residual Standard Error: 0.4581
```

```
> pLanc2 = loess.ancova(distance, cardbill, abcix) #about same result
> summary(pLanc2)
```

```
      Length Class  Mode
linear.fit  2      -none- numeric
smooth.fit 17      loess  list
```

```
> pLanc2
$linear.fit
```

```
          [,1]
(Intercept) 14605.635
group1      1127.854
```

```
$smooth.fit
```

```
Call:
```

```
loess(formula = lm.res ~ x, span = span1, degree = degree, family = family)
```

```
Number of Observations: 996
```

```
Equivalent Number of Parameters: 3.89
```

```
Residual Standard Error: 11160
```

`lindner {PSAgraphics}` R Documentation

Data on 996 initial Percutaneous Coronary Interventions (PCIs) performed in 1997 at the Lindner Center, Christ Hospital, Cincinnati.

Description

Data from an observational study of 996 patients receiving a PCI at Ohio Heart Health in 1997 and followed for at least 6 months by the staff of the Lindner Center. This is a landmark dataset in the literature on propensity score adjustment for treatment selection bias due to practice of evidence based medicine; patients receiving abciximab tended to be more severely diseased than those who did not receive a IIB/IIIa cascade blocker.

Usage `data(lindner)`

Format

A data frame with 996 observations on the following 10 variables, no NAs.

lifepres

Mean life years preserved due to survival for at least 6 months following PCI; numeric value of either 11.4 or 0. `cardbill`

Cardiac related costs incurred within 6 months of patient's initial PCI; numeric value in 1998 dollars; costs were truncated by death for the 26 patients with `lifepres == 0`. `abcix`

Numeric treatment selection indicator; 0 implies usual PCI care alone; 1 implies usual PCI care deliberately augmented by either planned or rescue treatment with abciximab. `stent`

Coronary stent deployment; numeric, with 1 meaning YES and 0 meaning NO. `height`

Height in centimeters; numeric integer from 108 to 196. `female`

Female gender; numeric, with 1 meaning YES and 0 meaning NO. `diabetic`

Diabetes mellitus diagnosis; numeric, with 1 meaning YES and 0 meaning NO. `acutemi`

Acute myocardial infarction within the previous 7 days; numeric, with 1 meaning YES and 0 meaning NO. `ejecfrac`

Left ejection fraction; numeric value from 0 percent to 90 percent. `veslproc`

Number of vessels involved in the patient's initial PCI procedure; numeric integer from 0 to 5.

Source Package USPS, by R. L. Obenchain.

Lindner Data (PCI)

Stat 266
week 2

```
##### Week 2 Computing Corner, Rogosa R-session
> library(PSAgraphics) > data(lindner) > attach(lindner)
> dim(lindner)
[1] 996 10
> head(lindner)
  lifepres cardbill  abcix  stent height female diabetic acutemi ejecfrac veslproc
1      0.0     14301      1      0     163      1      1      0      56      1
2     11.6     3563      1      0     168      0      0      0      56      1
3     11.6     4694      1      0     188      0      0      0      50      1
4     11.6     7366      1      0     175      0      1      0      50      1
5     11.6     8247      1      0     168      1      0      0      55      1
6     11.6     8319      1      0     178      0      0      0      50      1
```

mostly treated (evidence based medicine)

```
> table(abcix)
abcix
  0  1
298 698

##### look at outcomes
> tapply(cardbill, abcix, mean) > tapply(log(cardbill), abcix, mean) # analyses done in log scale
  0      1          0      1
14614.22 16126.68      9.398158 9.581579
```

```
> t.test(log(cardbill) ~ abcix) # treatment leads to higher bills?
Welch Two Sample t-test
data: log(cardbill) by abcix
t = -5.2317, df = 461, p-value = 2.554e-07
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval: -0.2523168 -0.1145249
mean in group 0 mean in group 1
  9.398158      9.581579
```

describe outcomes
men set aside

```
> table(lifepres, abcix)
```

```
      abcix
lifepres  0  1
  0      15 11
 11.6    283 687
```

```
> chisq.test(lifepres, abcix)
```

Pearson's Chi-squared test with Yates' continuity correction

```
data: lifepres and abcix
X-squared = 8.5077, df = 1, p-value = 0.003536
```

an association

```
> prop.table(table(lifepres, abcix), 2) # look at relative risk
```

```
      abcix
lifepres  0      1
  0      0.05033557 0.01575931
 11.6  0.94966443 0.98424069
```

```
> library(MatchIt) ## try full matching
```

```
> m2full = matchit(abcix ~ stent + height + female + diabetic + acutemi + ejecfrac + veslproc,
  data = lindner, method = "full")
```

```
> m2full.dat = match.data(m2full) > attach(m2full.dat) # get matched data (distance, subclass)
```

```
> boxplot(distance ~ abcix) # propensity score overlap; also overlapping histogram as in week 1
```

see figure

```
> summary(m2full) # check balance improvement
```

```
Call: matchit(formula = abcix ~ stent + height + female + diabetic +
  acutemi + ejecfrac + veslproc, data = lindner, method = "full")
```

```
Summary of balance for all data:
```

	Means Treated	Means Control	Mean Diff	eQQ Med	eQQ Mean	eQQ Max
distance	0.7265	0.6406	0.0859	0.0814	0.0852	0.1209
stent	0.7049	0.5839	0.1210	0.0000	0.1208	1.0000
height	171.4427	171.4463	-0.0036	0.0000	0.5638	20.0000
female	0.3309	0.3859	-0.0550	0.0000	0.0537	1.0000
diabetic	0.2049	0.2685	-0.0636	0.0000	0.0638	1.0000
acutemi	0.1791	0.0604	0.1187	0.0000	0.1174	1.0000
ejecfrac	50.4026	52.2886	-1.8860	1.0000	2.0503	20.0000
veslproc	1.4628	1.2047	0.2581	0.0000	0.2651	1.0000

```
Summary of balance for matched data:
```

	Means Treated	Means Control	Mean Diff	eQQ Med	eQQ Mean	eQQ Max
distance	0.7265	0.7262	0.0003	0.0068	0.0077	0.0798

stent	0.7049	0.7465	-0.0416	0.0000	0.0248	1.0000
height	171.4427	171.6093	-0.1666	0.0000	0.9548	15.0000
female	0.3309	0.3016	0.0293	0.0000	0.0244	1.0000
diabetic	0.2049	0.2210	-0.0162	0.0000	0.0068	1.0000
acutemi	0.1791	0.1605	0.0186	0.0000	0.0300	1.0000
ejecfrac	50.4026	50.9846	-0.5821	0.0000	0.9524	20.0000
veslproc	1.4628	1.4616	0.0012	0.0000	0.0324	1.0000

Percent Balance Improvement:

	Mean Diff.	eQQ Med	eQQ Mean	eQQ Max
distance	99.5977	91.6649	91.0174	33.9503
stent	65.6178	0.0000	79.4711	0.0000
height	-4508.6968	0.0000	-69.3633	25.0000
female	46.6926	0.0000	54.5550	0.0000
diabetic	74.5771	0.0000	89.3347	0.0000
acutemi	84.3069	0.0000	74.4571	0.0000
ejecfrac	69.1376	100.0000	53.5491	0.0000
veslproc	99.5373	0.0000	87.7782	0.0000

Control Treated

All	298	698
Matched	298	698
Unmatched	0	0
Discarded	0	0

```
> summary(m2full, standardize = T) # look at standardized mean diffs < .1
```

Summary of balance for matched data:

	Means Treated	Means Control	Std. Mean Diff.	eCDF Med	eCDF Mean	eCDF Max
distance	0.7265	0.7262	0.0027	0.0128	0.0123	0.0300
stent	0.7049	0.7465	-0.0911	0.0240	0.0240	0.0480
height	171.4427	171.6093	-0.0156	0.0128	0.0144	0.0444
female	0.3309	0.3016	0.0622	0.0094	0.0094	0.0188
diabetic	0.2049	0.2210	-0.0400	0.0044	0.0044	0.0088
acutemi	0.1791	0.1605	0.0485	0.0080	0.0080	0.0160
ejecfrac	50.4026	50.9846	-0.0559	0.0080	0.0109	0.0604
veslproc	1.4628	1.4616	0.0017	0.0044	0.0071	0.0196

```
> plot(summary(m2full, standardize = T)) # see plot, balance improvement
```

see picture

```
> head(m2full.dat)
```

```
> table(subclass) #267 subclasses (698 treated) mostly small #> table(subclass, abcix) breakdown
```

```
fivenum(table(subclass))
subclass
 3 154 266 147 190
 2  2  2  4  18
```

```
> library(lme4) compare outcomes over subclasses, log(cardbill) outcome
```

```
> mfullL.lmer = lmer(log(cardbill) ~ abcix + (1 + abcix|subclass), data = m2full.dat)
```

```
> summary(mfullL.lmer)
Formula: log(cardbill) ~ abcix + (1 + abcix | subclass) Data: m2full.dat
```

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
subclass	(Intercept)	0.09744	0.3122	
	abcix	0.06809	0.2609	-0.98
Residual		0.18034	0.4247	

Number of obs: 996, groups: subclass, 267

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	9.40576	0.03159	297.71
abcix	0.16440	0.03417	4.81

```
> confint(mfullL.lmer)
```

Computing profile confidence intervals ...

	2.5 %	97.5 %
.sig01	0.22608548	0.3882918
.sig02	-1.00000000	-0.8884949
.sig03	0.15108346	0.3579166
.sigma	0.40090446	0.4492170
(Intercept)	9.34378346	9.4678932

```

abcix      0.09694981  0.2327080
> exp(confint(mfullL.lmer))
(Intercept) 1.142756e+04 1.293760e+04
abcix      1.101805e+00 1.262013e+00

```

in g < vending machine

```

> # note mfull.lmer = lmer(cardbill ~ abcix + (1 + abcix|subclass), data = m2full.dat)
# gives non-sig result CI: abcix -497.3218 3009.1425564

```

```

##### treat lifepres as a 0,1 outcome
> mfull.glmer = glmer(as.factor(lifepres) ~ abcix + (1 + abcix|subclass),
                      family = binomial, data = m2full.dat)
> summary(mfull.glmer) # in log-odds metric
Generalized linear mixed model Family: binomial ( logit )
Random effects:
Groups Name Variance Std.Dev. Corr
subclass (Intercept) 0.003532 0.05943
abcix 42.902893 6.55003 1.00
Number of obs: 996, groups: subclass, 267

```

```

Fixed effects:
Estimate Std. Error z value Pr(>|z|)
(Intercept) 2.9389 0.2659 11.053 < 2e-16 ***
abcix 6.0279 0.8381 7.192 6.39e-13 ***

```

log odds

```

> exp(fixef(mfull.glmer)) #odds of survival
(Intercept) abcix
18.89487 414.82764

```

```

##### alternative, matching by regression interpolation, ancova vs blocking
> # propensity ancova, propensity score (distance) as covariate
> pancL = lm(log(cardbill) ~ abcix + distance, data = m2full.dat)
> summary(pancL)
Call: lm(formula = log(cardbill) ~ abcix + distance, data = m2full.dat)
Coefficients:

```

```

Estimate Std. Error t value Pr(>|t|)
(Intercept) 9.17752 0.07759 118.286 < 2e-16 ***
abcix 0.15384 0.03322 4.630 4.14e-06 ***
distance 0.34443 0.11378 3.027 0.00253 **
---

```

*ancova
straight-line*

```

> exp(confint(pancL)) # almost same result as fullmatch lmer
      2.5 %      97.5 %
(Intercept) 8310.388799 11268.542739
abcix 1.092686 1.244877
distance 1.128783 1.764223

```

```

> pLancL = loess.psa(log(cardbill), abcix, distance) # use smoothers rather than straight-line
> pLancL # similar result; loess.psa from PSAgraphics also generates plot
$ATE $se.wtd $CI95
[1] 0.1279487 [1] 0.04151743 [1] 0.04491388 0.21098361

```

see picture

loess

```

# compare a smoother ancova package fANCOVA
> install.packages("fANCOVA") > library(fANCOVA)
> pLanc2L = loess.ancova(distance, log(cardbill), abcix) #about same result
> pLanc2L
$linear.fit

```

ancova

```

[,1]
(Intercept) 9.3863321
group1 0.1522632

```

```

$smooth.fit
Call: loess(formula = lm.res ~ x, span = span1, degree = degree, family = family)

```

*full session posted.
more in RQ*